



Review Article

Association of Irritable Bowel Syndrome with Dementia: A GRADE-Assessed Systematic Review and Meta-Analysis

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ABSTRACT

Introduction: Numerous gastrointestinal (GI) diseases have been linked to an increased risk of dementia. Irritable bowel syndrome (IBS) is a prevalent GI disorder; however, existing studies investigating its association with dementia have reported mixed findings, and no prior meta-analysis has synthesized the available evidence.

Methods: A comprehensive search of MEDLINE (via PubMed), Scopus, and Web of Science was conducted in August 2025, with no language restrictions, to identify observational studies comparing the risk of all-cause dementia and dementia subtypes (e.g., Alzheimer's disease (AD)) in individuals with and without IBS. Pooled odds ratios (ORs) with 95% confidence intervals (CIs) were calculated using the DerSimonian–Laird random-effects model. The certainty of evidence was assessed using the GRADE approach.

Results: More than 750,000 participants were included from three cohort and two case-control studies. The pooled analysis showed that IBS was associated with significantly higher odds of all-cause dementia (OR = 1.22; 95% CI: 1.13–1.31; $p < 0.00001$; $I^2 = 26\%$), with the overall certainty of evidence rated as high. The association remained robust across all sensitivity analyses. IBS was also significantly associated with higher odds of AD (OR = 1.27; 95% CI: 1.08–1.50; $p = 0.005$; $I^2 = 64\%$), with the certainty of evidence rated as low.

Conclusion: Our findings consistently suggest that IBS is associated with higher rates of dementia outcomes; however, this association is derived from observational data and does not establish causality. Further well-designed prospective studies are needed to clarify the nature of this relationship.

1. Introduction

Dementia represents a major global health challenge, primarily affecting older adults. In 2021, the estimated prevalence was approximately 57 million people worldwide, with an annual incidence of around 10 million new cases [1]. It is characterized by progressive cognitive, functional, and behavioral decline that significantly impairs daily functioning [2]. Due to its prolonged course and the high level of care required, dementia imposes substantial financial burdens, with healthcare expenditures averaging 57% higher than those for other chronic conditions [3]. Therefore, investigating the modifiable risk factors through the preclinical phase of the disease is a global public health priority, especially given the lack of approved disease-modifying therapies [4].

A growing body of evidence highlights the role of gut health in cognitive function [5]. Disturbances in gastrointestinal (GI)

homeostasis have been linked to neurodegenerative processes, and some neurological disorders are thought to originate in the gut [6]. Studies suggest that gut inflammation, alterations in the gut microbiota, and changes in microbiota-derived metabolites may contribute to the pathogenesis of dementia [7]. These findings are supported by observational studies reporting significant associations between various GI disorders and the risk of developing dementia [8–10].

Among GI conditions, irritable bowel syndrome (IBS) is particularly common and impactful, affecting an estimated 5 – 10% of the global population [11]. It is characterized by recurrent abdominal pain or discomfort, altered visceral sensitivity, and changes in GI motility. While IBS is often diagnosed in individuals under 50, it can also affect older adults [12], making it a relevant condition to consider in dementia research.

According to recent observational research, there may be a connection between IBS and a higher risk of dementia. However, findings are inconsistent [13–17]. While multiple systematic reviews have examined GI disorders, namely liver diseases [18] and inflammatory bowel disease (IBD) [19, 20], in relation to dementia, no comprehensive meta-analysis has focused on IBS. Given its high prevalence and distinct pathophysiology, IBS warrants separate investigation.

Therefore, this study aims to conduct a systematic review and meta-analysis to assess the association between IBS and dementia risk, including both incident risk estimates from cohort studies and

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exposure-outcome associations from case-control studies, and to explore potential differences by dementia subtype.

2. Methods

The protocol for this systematic review and meta-analysis was registered a priori on the International Prospective Register of Systematic Reviews (PROSPERO; ID: CRD420251124693). Reporting adhered to the Preferred Reporting Items for Systematic Reviews and Meta-Analyses (PRISMA) guidelines, and all steps were conducted following guidelines stated in the Cochrane Collaboration Handbook [21].

2.1. Inclusion and exclusion criteria

Eligibility criteria were structured using the PICOS framework:

- Population: Adults from any population, with or without IBS at baseline.
- Intervention/Exposure: Presence or diagnosis of IBS, defined by clinical diagnosis, ICD codes, or validated criteria (e.g., Rome criteria).
- Comparator: Individuals without IBS.
- Outcome: Diagnosis of dementia (any type: all-cause, Alzheimer's disease (AD), vascular dementia (VaD), etc.), diagnosed clinically or via registry data.
- Study design: Observational studies (cohort or case-control studies) reporting on the association between IBS and dementia risk.

Only studies employing valid diagnostic methods for both IBS (exposure) and dementia (outcome) were eligible. No restrictions were placed on publication year or language. We excluded animal studies, non-comparative designs, reviews, editorials, and commentaries.

2.2. Study selection

A systematic literature search was conducted on August 5, 2025, across three electronic databases: PubMed, Scopus, and Web of Science. The search strategy contained terms related to irritable bowel syndrome (e.g., "IBS," "functional bowel disorder") and dementia (e.g., "Alzheimer's disease," "cognitive impairment"). The full search syntax and fields for each database are provided in (Supplementary Table 1).

All identified records were imported into Rayyan software for screening. Two independent reviewers first screened titles and abstracts for relevance. Studies that appeared eligible were then subjected to full-text review against our inclusion criteria. In addition to the database search, a manual review of the reference lists of included studies was performed to capture any additional relevant evidence.

2.3. Data extraction and assessment of risk of bias

A standardized online form was developed to extract data on study characteristics, population details, outcome measures, adjustment for confounders, and risk-of-bias assessments. Data extraction was performed independently by two investigators. Web-based tools for data conversion, such as the Meta-Analysis Accelerator software [22], were used when necessary. For instance, in one study that did not report the mean age and standard deviation (SD) for the combined AD and control groups, we applied this tool to calculate the pooled mean and SD from the subgroup values.

The risk of bias for included studies was assessed using the Newcastle-Ottawa Scale (NOS) [23]. The NOS evaluates three

domains: (1) selection of study participants (maximum 4 points), (2) comparability of groups (maximum 2 points), and (3) outcome assessment (cohort) or exposure ascertainment (case – control) (maximum 3 points). Scores range from 0 to 9, with ≥ 7 considered low risk of bias.

2.4. Statistical analysis

2.4.1. Statistical analysis plan

Our analysis examined the association between IBS diagnosis and dementia outcomes (all-cause and subtypes). This association is measured and reported in individual studies as an odds ratio (OR) or hazard ratio, with their corresponding 95% confidence intervals (CI). Given the relatively low incidence of dementia, these measures were considered approximately equivalent for pooling, an approach supported by established epidemiological principles and widely adopted in high-quality meta-analyses [24, 25].

When multiple effect estimates with different levels of confounder adjustment were reported, the most fully adjusted estimate was extracted to better approximate the true association while minimizing residual confounding. We also performed subgroup analyses stratified by the level of confounder adjustment to explore its potential impact on the pooled estimates.

Main meta-analyses were performed using RevMan software (Cochrane Collaboration). Extracted ORs with their 95% CIs from each study were entered into RevMan's built-in calculator, which automatically derives the log OR and its standard error. A random-effects model based on the DerSimonian – Laird (DL) method was employed to account for expected within- and between-study variability.

Given the small number of included studies and the anticipated substantial heterogeneity, the DL method's tendency to underestimate between-study variance in such settings was a recognized limitation [26]. To assess the robustness of the primary pooled estimates, a sensitivity analysis was therefore conducted using the Restricted Maximum Likelihood (REML) estimator for τ^2 , implemented in R (online version of RStudio; meta and metafor packages) [26]. REML provides a more conservative variance estimate than DL and was used to assess whether the main inference remained stable under a stricter model of heterogeneity. A p-value < 0.05 was considered statistically significant throughout.

2.4.2. Heterogeneity, sensitivity analyses, and subgroup analyses

Heterogeneity was assessed visually with forest plots and statistically using the chi-square test ($p < 0.10$ considered significant) and the I^2 statistic [25]. Publication bias was not formally assessed with funnel plots, consistent with recommendations by Sterne et al., as fewer than 10 studies were included [27].

Sensitivity analyses were performed using the leave-one-out method to assess the influence of individual studies on overall effect estimates and determine the robustness of findings.

Planned subgroup analyses were performed to investigate potential sources of heterogeneity, including study design, geographic region, and depth of confounder adjustment. For confounder adjustment, studies were categorized as minimal, moderate, or extensive based on the variables included. Minimal adjustment included only demographic factors; moderate adjustment included demographics plus clinical comorbidities (e.g., diabetes, stroke); and extensive adjustment included genetic factors (e.g., APOE $\epsilon 4$ allele).

2.4.3. Certainty of evidence assessment

The certainty of evidence was evaluated using the Grading of Recommendations Assessment, Development and Evaluation (GRADE) approach via GRADEpro software [28]. For observational studies evaluating prognostic associations (i.e., the association between IBS and dementia risk), GRADE begins with a high certainty rating [29]. It may be downgraded based on risk of bias, indirectness of evidence, inconsistency (heterogeneity), imprecision in effect estimates, and publication bias (when assessable). Upgrading was considered in cases of large effect sizes or evidence of dose-response relationships. The final certainty rating was categorized as high, moderate, low, or very low.

2.5. Study-specific methodological considerations

One included record (Konings et al., 2023) described two distinct studies that used the US-based TriNetX database [14]. First, the authors conducted a case-control study to examine the association between various diseases (including AD) and a prior diagnosis of IBS. To validate these findings, they performed a complementary cohort study within an IBS-diagnosed population from the same database, investigating the subsequent risk of developing AD. Because both analyses draw from the same database (TriNetX) during the same time frame (2005 to 2021), a single patient who has both an IBS diagnosis and a subsequent AD diagnosis would likely be included in both analyses. Including both estimates in the meta-analysis would therefore introduce double-counting of participants, violating the assumption of independent study samples and potentially inflating the pooled effect size [21]. Accordingly, only the primary (case-control) analysis was included and is referred to throughout as Konings et al. (2023); the cohort analysis was excluded from all pooled estimates.

Konings et al. (2023) reported the odds of developing AD (specifically AD rather than all-cause dementia) in relation to three distinct IBS subtypes (IBS with constipation, IBS with diarrhea, and IBS without diarrhea). Still, they did not provide a single combined effect estimate [14]. However, the authors did not clarify whether these subgroups were mutually exclusive or whether a hierarchical classification was applied to prevent overlap. Since IBS subtypes coded in routine data are often not mutually exclusive over time (patients can carry constipation and diarrhea codes at different encounters), combining these estimates in an internal meta-analysis could introduce unit-of-analysis errors and double counting. Therefore, we adopted a single-estimate approach and selected the estimate for "IBS without diarrhea" (ICD-10: K58.9) as the representative effect. We chose this specific endpoint because the authors define it as a synonym for "IBS not otherwise specified," and it constitutes the largest patient cohort in their study ($N = 1,032$ vs. $N = 51$ for IBS-Constipation and $N = 143$ for IBS-Diarrhea).

Another study (Tung et al. 2022) investigated the association between functional gastrointestinal disorders (FGIDs) and all-cause dementia without reporting IBS-specific data separately [15]. Nevertheless, this study was included because IBS is the most frequent FGID [30–34] and excluding it would omit relevant evidence.

To ensure strict comparability of both exposure and outcome definitions, the primary meta-analysis was restricted to studies evaluating all-cause dementia following an exclusive diagnosis of IBS. Therefore, Tung et al. 2022 (FGID as exposure) and Konings 2023 (AD as outcome only) were not included in the primary pooled analysis of the all-cause dementia outcome. A broader secondary analysis was subsequently conducted to provide a more comprehensive overview of dementia outcomes associated with IBS

and related FGIDs by incorporating all five originally identified studies.

3. Results

3.1. Search results

The initial search yielded 467 records. After removing 61 duplicates, 406 records remained for title and abstract screening. Of these, 392 records were excluded for failing to meet the eligibility criteria. A total of 13 studies underwent full-text review, and eight were excluded for two primary reasons: the exposure of interest was not IBS [35, 36], or the studies did not report the risk of dementia or one of its subtypes [37–39] (**Figure 1**). The manual search (citation analysis) identified no additional studies. Ultimately, five studies met all eligibility criteria and were included in the final systematic review and meta-analysis [13–17].

3.2. Study characteristics

This systematic review included five studies: three cohort studies [13, 16, 17] and two case-control studies [14, 15]. Together, the studies had 759,365 participants, with sample sizes ranging from 6,600 to 458,181. The studies were conducted in three main countries: the United States (US), the United Kingdom (UK), and Taiwan. The male population in the studies ranged from 47.3% to 90.2%. General characteristics of the included studies and populations are summarized in (**Table 1**).

The inclusion criteria varied slightly by study design. In case-control studies (Konings et al. 2023 and Tung et al. 2022), participants generally required a pre-existing diagnosis of dementia. In contrast, the cohort studies (Chen et al. 2016, Yeh et al. 2018, and Yuan et al. 2024) included participants with a newly diagnosed IBS. In most cases, diagnoses were established using the International Classification of Diseases (ICD)-9 or ICD-10 codes, although some studies also incorporated physician-confirmed diagnoses. Detailed inclusion and exclusion criteria, study methodologies, and dementia subtypes assessed are presented in (**Table 2**).

3.3. Pooled risk of dementia

Some studies reported the risk of all-cause dementia alongside specific subtypes, including AD, VaD, or non-AD dementia, while others reported all-cause dementia as the sole outcome. The extracted effect estimates for each reported dementia outcome from individual studies are summarized in (**Table 3**).

The primary restricted analysis, limited to IBS as the exposure and all-cause dementia as the outcome ($n = 3$ studies), demonstrated that individuals with IBS had 22% higher odds of developing all-cause dementia compared with controls (OR = 1.22; 95% CI: 1.13 – 1.31; $p < 0.00001$, $I^2 = 26\%$) (**Figure 2**).

The broader secondary analysis incorporating all five studies (see Section 2.5) demonstrated that individuals with IBS/FGID had 38% higher odds of developing dementia outcomes compared with controls (OR = 1.38; 95% CI: 1.16 – 1.64; $p = 0.0002$). This analysis showed significant heterogeneity ($I^2 = 86\%$, $p < 0.0001$) (**Figure 3**).

3.3.1. Sensitivity analyses

To explore potential sources of heterogeneity and assess the robustness of the results, we performed a leave-one-out sensitivity analysis in both the primary and secondary analyses. Results remained consistent across all study exclusions in both analyses (**Table 4**) and (**Table 5**); (**Supplementary Figures 1 and 2**).

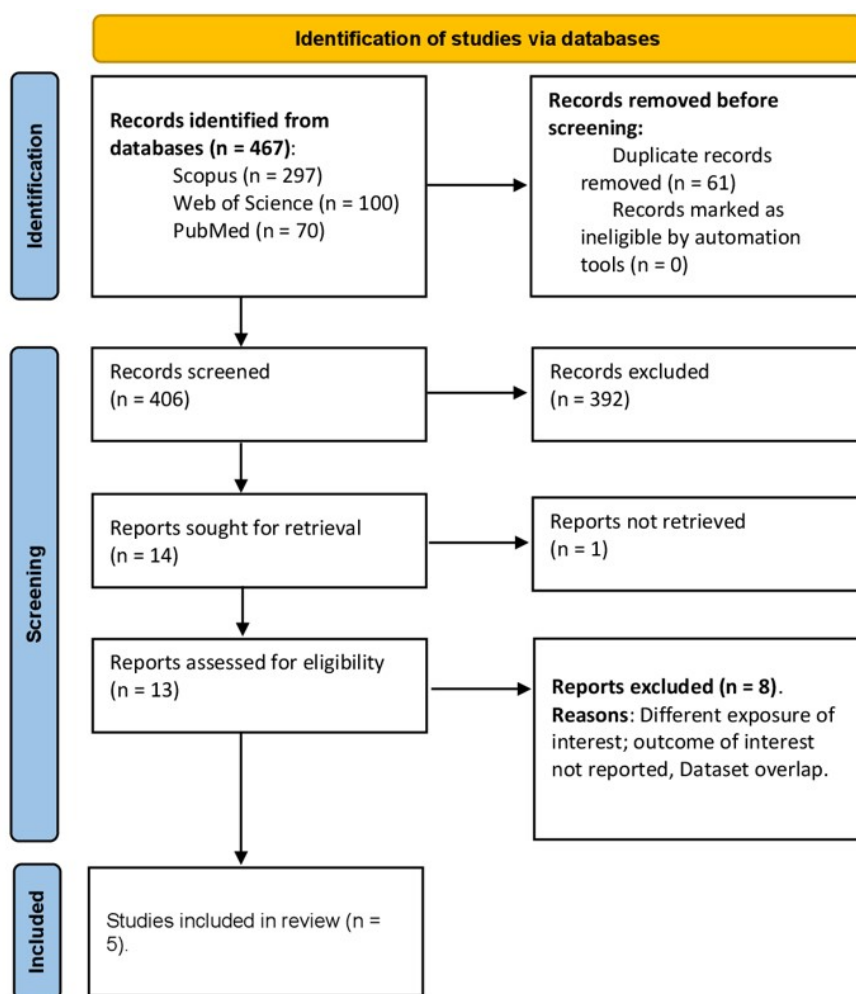


Figure 1: PRISMA flow diagram of the study selection process.

Table 1: General characteristics of included studies and their populations [13–17]

Study ID	Study design	Data source	Location	Study duration (years)	Total sample size	Mean age (SD)	Male n (%)
Chen 2016	Cohort	NHIRD	Taiwan, Asia	2000–2011 (11)	161,490	51.1 (16.6)	76,520 (47.3)
Konings 2023	Case-control	TriNetX	US, North America	2005–2021 (16.5)	43,670	71.65 (8.53)	24,233 (55.5)
Tung 2022	Case-control	TPMI	Taiwan, Asia	2013–2021 (8)	6,600	73.0 (9.3)	5,950 (90.2)
Yeh 2018	Cohort	NHIRD	Taiwan, Asia	2000–2015 (16)	89,424	55.1 (18.3)	47,920 (53.6)
Yuan 2024	Cohort	UK Biobank	UK, Europe	2006–2021 (15)	458,181	56.7 (2.1)	209,373 (45.7)

NHIRD, National Health Insurance Research Database; TPMI, Taiwan Precision Medicine Initiative; US, United States; UK, United Kingdom; SD, standard deviation

3.3.2. Exploratory subgroup analyses

Subgroup analyses were conducted to explore key influencing factors by stratifying studies according to study design, geographical area, and depth of confounder adjustment (**Table 6**); (**Supplementary Figures 3–5**). These analyses were applied to the broader secondary dementia pool rather than the primary restricted pool, as the larger dataset reduced the risk of yielding single-study subgroups – an outcome that would limit any meaningful contribution to characterizing the IBS – dementia association.

3.4. Pooled risk of Alzheimer's disease

Among dementia subtypes, AD was the only outcome reported in more than one study (**Table 2**), column "Type of dementia assessed"; (**Table 3**). In the pooled analysis of three studies (Chen et al. 2016; Konings et al. 2023; Yuan 2024), IBS was significantly associated with increased odds of AD (OR = 1.27; 95% CI: 1.08 – 1.50; $p = 0.005$), with moderate heterogeneity ($I^2 = 64\%$, $p = 0.06$) (**Figure 4**).

In the leave-one-out sensitivity analysis, the direction of the association remained consistent across all iterations; however, statistical significance was lost upon individual exclusion of Konings et al.

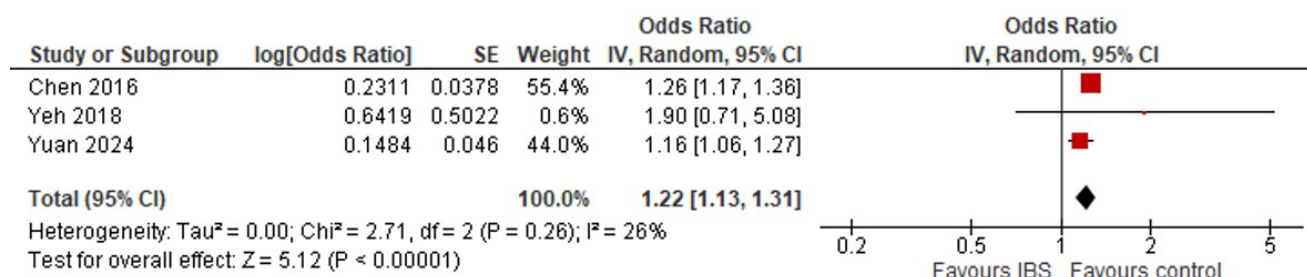


Figure 2: Meta-analysis of association between exclusive IBS and all-cause dementia.

(2023) and Yuan et al. (2024), indicating that the pooled estimate is sensitive to these two studies.

3.5. Meta-analysis using the REML approach

To assess the robustness of the primary pooled estimates against a more conservative heterogeneity model, all meta-analyses were re-run using the REML approach. A comparative summary of the DL and REML estimates is presented in (Table 7), with the corresponding REML forest plots provided in (Supplementary Figures 6 - 8).

Consistent with the statistical properties of the REML estimator, this method produced wider CIs and higher p-values across the included analyses. Despite these adjustments for between-study variance, neither the magnitude of the effect estimates nor the overall statistical significance was materially altered (Table 7).

3.6. GRADE assessment

We applied the GRADE methodology to evaluate the certainty of evidence for the outcomes of this review (Table 8). For the primary analysis of all-cause dementia, the certainty of evidence was rated as high. For the secondary analysis of dementia outcomes, certainty was rated as low, with downgrading applied in the domains of inconsistency/heterogeneity and indirectness. Although the observational nature of the included studies is inherently associated with greater heterogeneity in meta-analyses [40], and some authors consider these insufficient grounds for downgrading, we judged the degree of heterogeneity in this analysis to be substantial enough to warrant downgrading ($I^2 = 83\%$; $p < 0.0001$).

For the AD outcome, certainty was similarly rated as low, with downgrading applied in the domains of inconsistency and indirectness. The downgrading for indirectness was driven by the fact that, in one of the three contributing studies, AD was not the primary outcome of interest; the study was primarily designed to investigate Parkinson's disease [14].

3.7. Risk of bias assessment

The methodological quality of the included studies was evaluated using NOS (Table 2). Scores ranged from 8 to 9 out of 9, indicating high quality and adequate fulfillment of most NOS domains.

4. Discussion

4.1. Summary of main findings

This systematic review and meta-analysis examined the association between IBS and dementia across three pre-specified analyses. In the primary analysis, restricted to cohort studies with IBS-exclusive exposure and all-cause dementia as the outcome, IBS was associated

with a 22% increase in the odds of developing all-cause dementia (OR = 1.22; 95% CI: 1.13 – 1.31), with low heterogeneity ($I^2 = 26\%$) and high-grade certainty of evidence. The secondary analysis, which incorporated all five studies including those with broader FGID exposure and mixed dementia outcomes, yielded a larger point estimate (OR = 1.38; 95% CI: 1.16 – 1.64), though this was accompanied by substantial heterogeneity ($I^2 = 86\%$) and was rated as low certainty by GRADE. In the AD-specific analysis, IBS was associated with a 27% increase in the odds of AD (OR = 1.27; 95% CI: 1.08 – 1.50), with moderate heterogeneity ($I^2 = 64\%$) and low-grade certainty.

The leave-one-out sensitivity analyses demonstrated that the primary and secondary pooled estimates were robust to the removal of any individual study, with the direction and magnitude of the association remaining consistent across all iterations. In contrast, the AD-specific estimate was sensitive to the individual exclusion of Konings et al. (2023) and Yuan et al. (2024), with statistical significance not maintained upon their removal, warranting cautious interpretation of this subtype finding.

Exploratory subgroup analyses applied to the secondary dementia pool revealed notable patterns. Cohort studies demonstrated smaller, statistically significant, and more consistent effects, whereas case-control studies showed larger but non-significant and highly heterogeneous estimates. This discrepancy may be attributed to inherent methodological limitations and the difficulty of establishing temporal relationships in case-control designs. Furthermore, an age-gap analysis revealed that case-control cohorts were approximately 15 – 20 years older than their longitudinal counterparts (Table 1). Given that age is the primary driver of dementia incidence, this demographic shift likely amplifies the variance in effect sizes [41, 42].

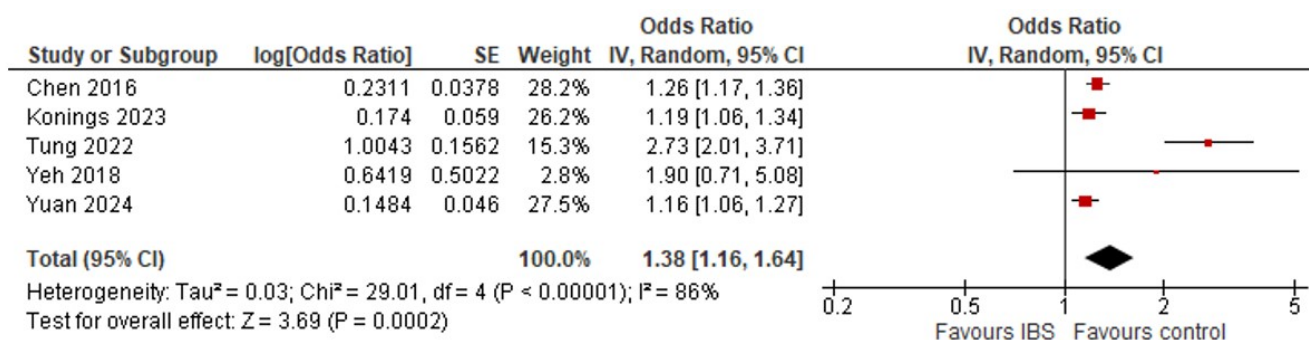
Moreover, studies conducted in Asian populations, comprising three studies, yielded a non-significant pooled estimate, suggesting that the IBS – dementia association may not be uniform across ethnicities and warrants further investigation in Asian cohorts. In contrast, the American and European subgroups each comprised a single study, precluding any meaningful regional interpretation for these populations. Similarly, the subgroup of studies employing extensive confounder adjustment – including genetic factors such as the APOE $\epsilon 4$ allele – produced a non-significant result, in contrast to the minimal and moderate adjustment subgroups, raising the possibility that part of the observed IBS – dementia association may be mediated or confounded by genetic susceptibility rather than IBS-specific pathophysiology.

Finally, re-running all pooled analyses using the REML estimator, a more conservative approach to between-study variance, produced

Table 2: Methodological characteristics of included studies [13–17]

Study ID	Inclusion criteria	Exclusion criteria	Diagnosis of IBS	Subjects with IBS n (%)	Type of dementia assessed	Diagnosis of dementia	Dementia cases n (%)	NOS score
Chen 2016	Patients aged ≥ 20 years with newly diagnosed IBS (2000–2011) identified from the LHID2000 database. A non-IBS cohort was randomly selected from individuals without IBS during the same period and matched.	Patients with a prior history of dementia or missing medical information.	ICD-9-CM codes	32,298 (20.0)	All-cause, AD, non-AD dementia	ICD-9-CM codes	4,062 (2.5)	9
Konings 2023	First diagnosis of AD between ages 50–90, with ≥ 2 years of ambulatory visits before diagnosis (between Jan 1, 2005, and July 1, 2021).	Patients with a diagnosis of dementia before their first AD diagnosis.	ICD-10 codes	NA	AD	ICD-10 codes	19,046 (43.6)	9
Tung 2022	Dementia cases diagnosed between Jan 1, 2013, and Apr 30, 2021; controls without dementia up to Apr 30, 2021; age- and sex-matched at a 1:10 ratio.	Patients who did not consent to data inclusion.	NA	282 (4.2)	All-cause dementia	ICD-9 codes	600 (9.1)	9
Yeh 2018	Adults with ≥ 3 outpatient visits for IBS within 1 year.	Patients < 20 years or with psychiatric disorders diagnosed before 2000 or before the first IBS visit.	Rome criteria	22,356 (25.0)	All-cause dementia	Diagnosed by psychiatrists	672 (0.7)	8
Yuan 2024	Adults without baseline dementia and with available genetic information.	Patients with dementia at baseline, dementia within the first year of follow-up, or missing genetic data.	Diagnostic codes from inpatient, primary care, registries, self-report	480 (0.1)	All-cause dementia, AD, VaD	ICD-9/10 codes	6,415 (1.4)	9

IBS, irritable bowel syndrome; AD, Alzheimer's disease; VaD, vascular dementia; ICD-9-CM, International Classification of Diseases, 9th Revision, Clinical Modification; ICD-10, International Classification of Diseases, 10th Revision; NOS, Newcastle–Ottawa Scale; NA, not available

**Figure 3:** Meta-analysis of the association between IBS/FGID and dementia outcomes.

wider confidence intervals and attenuated p-values, as expected. Despite this, neither the direction nor the overall statistical significance of the primary estimates was materially altered, supporting the robustness of the findings across heterogeneity modeling approaches.

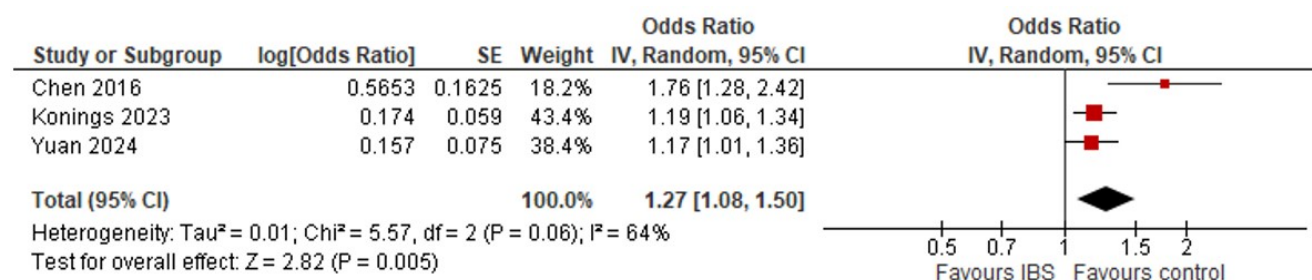
4.2. Biological plausibility

Several biological mechanisms may underlie the observed association between IBS and dementia outcomes (all-cause and AD) in our analyses, with the link largely involving the brain – gut axis and shared pathological pathways [43]. In IBS, symptoms

Table 3: Summary of extracted effect estimates, dementia outcomes, and adjustments by study [13–17]

Study ID	Fully adjusted effect estimate	95% CI (LL)	95% CI (UL)	Variables adjusted for
Chen 2016	All-cause dementia: HR = 1.26*	1.17	1.35	Age, sex, diabetes, hypertension, stroke, CAD, head injury, depression, epilepsy
	AD dementia: HR = 1.76#	1.28	2.43	
	Non-AD dementia: HR = 1.24	1.15	1.33	
Konings 2023	AD dementia: OR = 1.19*#	1.06	1.34	Age, sex, race, ethnicity
Tung 2022	FGID and all-cause dementia: OR = 2.73*	2.01	3.71	APOE ϵ 4 allele, CVA, sleep, FGIDs
Yeh 2018	All-cause dementia: HR = 1.90*	0.71	2.57	Age, sex, education, marital status, comorbidities (CCI), season, residence, insurance
Yuan 2024	All-cause dementia: HR = 1.16*	1.06	1.27	Age, sex, deprivation index, education, BMI, activity, diet, smoking, alcohol, HTN, stroke, depression, polygenic risk score
	AD dementia: HR = 1.17#	1.01	1.35	
	VaD dementia: HR = 1.38	1.13	1.70	

AD, Alzheimer's disease; IBS, irritable bowel syndrome; VaD, vascular dementia; HR, hazard ratio; OR, odds ratio; CI, confidence interval; LL, lower limit; UL, upper limit; FGID, functional gastrointestinal disorder; CAD, coronary artery disease; APOE, apolipoprotein E; CVA, cerebrovascular accident; FGIDs, functional gastrointestinal disorders; CCI, Charlson comorbidity index; BMI, body mass index; HTN, hypertension; * Estimates used in the dementia outcomes analysis; # Estimates used in the Alzheimer's disease analysis

**Figure 4:** Meta-analysis of the association between irritable bowel syndrome and Alzheimer's disease.**Table 4:** Leave-one-out sensitivity analyses for the primary meta-analysis (exclusive IBS and all-cause dementia)

Excluded study	OR (95% CI)	p-value	Chi-squared p-value, I^2
Chen 2016	1.16 (1.06; 1.27)	0.0009	0.33, 0%
Yeh 2018	1.21 (1.12; 1.32)	< 0.00001	0.16, 48%
Yuan 2024	1.26 (1.17; 1.36)	< 0.00001	0.41, 0%

Significant test p-values are bold.

are driven by abnormal signaling within the bidirectional gut – brain axis, which integrates communication between the enteric nervous system (ENS), the central nervous system (CNS), and the gut microbiota [44]. Alterations in either the CNS or the GI tract can disrupt the gut microbiota, leading to dysbiosis and gut inflammation. These changes can signal back to the brain through vagal anti-inflammatory pathways [45], potentially contributing to altered brain function and neurodegeneration. In addition, IBS increases intestinal permeability, allowing bacterial products such

Table 5: Leave-one-out sensitivity analyses for the secondary broader meta-analysis

Excluded study	OR (95% CI)	p-value	Chi-squared p-value, I^2
Chen 2016	1.50 (1.13; 2.00)	0.006	< 0.00001, 90%
Konings 2023	1.50 (1.18; 1.90)	0.001	< 0.00001, 89%
Tung 2022	1.21 (1.15; 1.28)	< 0.00001	0.41, 0%
Yeh 2018	1.37 (1.15; 1.63)	0.0004	< 0.00001, 89%
Yuan 2024	1.53 (1.18; 1.98)	0.001	< 0.0001, 88%

Significant test p-values are bold.

as lipopolysaccharides (LPS) to enter circulation, thereby promoting neuroinflammation and increasing dementia risk [46].

Table 6: The results of subgroup analyses based on key variables

	Number	OR (95% CI)	p-value	Chi-squared p-value, I ²
All studies	5	1.38 (1.16; 1.64)	0.0002	< 0.0001, 86%
Study design				
Cohort	3	1.22 (1.13; 1.31)	< 0.00001	0.26, 26%
Case-control	2	1.78 (0.79; 4.01)	0.16	< 0.0001, 96%
Geographical region				
Asia (Taiwan)	3	1.84 (0.97; 3.48)	0.06	< 0.00001, 92%
America (USA)	1	1.19 (1.06; 1.34)	0.003	Not applicable
Europe (UK)	1	1.16 (1.06; 1.27)	0.001	Not applicable
Depth of adjustment for confounders				
Minimal Adjustment	1	1.19 (1.06; 1.34)	0.003	Not applicable
Moderate Adjustment	2	1.26 (1.17; 1.36)	< 0.00001	0.41, 0%
Extensive Adjustment	2	1.76 (0.76; 4.06)	0.19	< 0.00001, 96%

Significant test p-values are bold.

Table 7: Comparison of DerSimonian – Laird and Restricted Maximum Likelihood estimates across meta-analyses

Analysis	Model	OR (95% CI)	p-value
Exclusive IBS and all-cause dementia	DL	1.22 (1.13; 1.31)	< 0.00001
	REML	1.22 (1.13; 1.32)	< 0.0001
IBS/FGID and dementia outcomes	DL	1.38 (1.16; 1.64)	0.0002
	REML	1.49 (1.08; 2.06)	0.0165
IBS and AD	DL	1.27 (1.08; 1.50)	0.005
	REML	1.29 (1.05; 1.57)	0.014

OR, odds ratio; CI, confidence interval; DL, DerSimonian–Laird; REML, restricted maximum-likelihood; IBS, irritable bowel syndrome; FGID, functional gastrointestinal disorder; AD, Alzheimer's disease.

Immune activation further supports this pathway: binding of LPS to Toll-like receptor 4 (TLR4) activates the NF- κ B pathway in CNS cells [47], triggering the release of proinflammatory cytokines such as IL-6, IL-1 β , and TNF- α [48]. These cytokines not only influence enteric nerve activity but also stimulate the brain's immune system, thereby promoting neuroinflammation, a key process in neurodegeneration and cognitive decline [49]. Cytokines have also been independently associated with increased dementia risk [46]. Moreover, studies [50, 51] have shown strong pathological links between microbiome-derived neurotoxins and impaired synaptic signaling in AD. This process, combined with neuroinflammation, accelerates amyloid- β deposition in memory-related regions such as the hippocampus and cerebral cortex [52].

Psychiatric and behavioral comorbidities provide additional explanatory pathways. IBS is commonly associated with depression, anxiety, and sleep disturbances [53], all established risk factors for dementia [54]. IBS patients also exhibit alterations in neurotransmitters such as serotonin, dopamine, GABA, and brain-derived neurotrophic factor (BDNF) [46], which regulate gut motility, visceral sensitivity, and brain function [48]. Finally, shared vascular and metabolic risk

factors, including hypertension, diabetes, and obesity, may contribute to both GI dysfunction and cognitive decline [55].

Taken together, these multifaceted pathways suggest that the observed association likely reflects a cumulative, synergistic interplay between gut-derived neurotoxicity, systemic metabolic vulnerability, and the long-term cognitive burden of associated psychiatric comorbidities.

4.3. Limitations and strengths

This review has several limitations, many of which reflect concerns inherent to the existing evidence base on IBS and dementia risk. Potential surveillance or detection bias is a concern, as most studies relied on healthcare databases in which individuals with IBS are more likely to have frequent medical visits than the general population [13, 56]. This increases the likelihood of cognitive assessment and earlier or more frequent dementia diagnosis, potentially inflating the observed association independent of a true biological link. Although studies adjusted for some confounders, residual bias related to healthcare access and diagnostic intensity cannot be excluded. Future research using designs that better control for healthcare utilization, such as matched cohorts with similar contact frequency or prospective population-based studies, is needed to clarify this effect.

Residual confounding also remains a concern. While most studies adjusted for basic demographic and clinical variables, several shared risk factors between IBS and dementia were not consistently controlled for (**Table 3**). These include psychiatric comorbidities (e.g., depression and anxiety), sleep disturbances, medication burden (particularly psychotropic and anticholinergic drugs), and healthcare-seeking behavior. Given that these factors are strongly associated with both IBS and cognitive decline, they may partially explain the observed association [53, 57–60]. Findings from exploratory subgroup analyses further support this interpretation, as studies with limited or moderate adjustment reported significant associations, whereas more extensively adjusted studies showed attenuated effects. Therefore, unmeasured or residual confounding cannot be ruled out, highlighting the need for future studies with more comprehensive adjustment and advanced analytical approaches, such as propensity

Table 8: Certainty of evidence assessment

Analysis	Number of studies	Study design	Risk of bias	Inconsistency	Indirectness	Imprecision	Other considerations*	Overall certainty
Exclusive IBS and all-cause dementia	3	Observational	not serious	not serious	not serious	not serious	none	High
IBS/FGID and dementia outcomes	5	Observational	not serious	serious	serious	not serious	none	⊕ ⊕ ○○
								Low
IBS and AD	3	Observational	not serious	serious	serious	not serious	none	⊕ ⊕ ○○
								Low

*Other considerations: publication bias, large effect, plausible confounding, dose-response gradient

score matching or causal inference methods, to better isolate the independent effect of IBS on dementia risk.

Reverse causation may also contribute to the observed association. Early neurodegenerative changes during the prodromal phase of dementia may influence GI function through autonomic dysregulation or behavioral changes, leading to IBS-like symptoms before formal dementia diagnosis and increased healthcare utilization, thereby raising the likelihood of an IBS diagnosis [61–63]. Consequently, the observed association may partly reflect early neurodegeneration rather than a causal effect of IBS on dementia risk. This issue is particularly relevant in retrospective database studies. It highlights the need for longitudinal designs with clear temporal separation between IBS onset and dementia diagnosis, as well as the application of lag-time analyses.

Variability in exposures and outcomes across included studies is an additional limitation. Although we restricted the primary analysis to IBS-exclusive exposure and all-cause dementia as the outcome – and the AD-specific analysis to IBS and AD only – this resulted in only three studies contributing to each analysis, constituting a limited evidence base.

Furthermore, the primary objective of Konings et al. (2023) was to investigate the association between IBS and Parkinson's disease, with AD data reported as a secondary outcome [14]. Although this limitation was formally accounted for through downgrading of the certainty of evidence for indirectness in the AD-specific GRADE assessment [64], study power and reliability for the AD outcome remain inherently lower than in studies specifically designed to investigate it. Finally, three of the five included studies were conducted in Taiwan, which may limit generalizability to other populations.

This review also has several notable strengths. To our knowledge, it is the first meta-analysis to comprehensively synthesize evidence on the association between IBS and dementia. The large, pooled sample size (>750,000 participants) conferred substantial statistical power, and the included studies were of high methodological quality, with NOS scores of 8 to 9 out of 9, indicating low risk of bias. The review was conducted in accordance with the principles of the Cochrane Handbook and adhered to PRISMA reporting guidelines. Methodological rigor was further ensured through explicit handling of participant overlap to avoid double-counting (Section 2.5), referencing the excluded studies in the full-text screening and transparent reporting of all study-level inputs (Table 3) to ensure reproducibility, and the use of two complementary meta-analytic

models (DL and REML) to assess the stability of pooled estimates across heterogeneity modeling approaches.

4.4. Position in the relevant literature

The gut – brain axis seems not to be only a preclinical theory demonstrated in vitro or in animal models. Especially in neurodegeneration, many GI diseases have been linked to an increased risk of various neurodegenerative disorders, particularly AD. Among these GI disorders, IBD and NAFLD (non-alcoholic fatty liver disease) have attracted major interest, with a large number of observational studies exploring their association with AD and all-cause dementia and several meta-analyses already establishing this evidence [18–20]. Other GI diseases are now entering this line of investigation in observational studies but still lack strong and comprehensive meta-analytic evidence; for example, peptic ulcers, gastritis, and colorectal cancer [65–68].

4.5. Implications

This systematic review and meta-analysis suggest a potential association between IBS and higher odds of all-cause dementia and AD. However, the evidence derives from a limited number of observational studies. It is subject to substantial heterogeneity and potential biases, including surveillance bias, residual confounding, and reverse causation, which may partially or wholly account for the observed relationship.

Accordingly, IBS should not be construed as an independent causal risk factor for dementia based on current evidence. Rather, these hypothesis-generating findings highlight the importance of vigilance in patients with IBS who carry additional shared risk factors, such as depression, sleep disturbance, or extensive medication burden, where cumulative risk may be clinically meaningful. Current data do not justify changes in dementia-specific clinical management for IBS patients; however, they reinforce the value of comprehensive, longitudinal patient assessment in this population.

Future research should prioritize clarifying the direction and underlying mechanisms of the IBS-dementia association through well-designed prospective studies with clearly defined exposures and outcomes, adequate temporal separation between IBS onset and dementia diagnosis, lag-time analyses, and rigorous control for confounding. Separate evaluation of dementia subtypes, including AD, VaD, and non-AD dementia, is also needed, as pooling across subtypes may obscure clinically distinct pathophysiological pathways [69]. Mechanistic studies exploring the gut-brain axis, gut microbiota composition, and systemic neuroinflammation may further elucidate the biological basis of any true association.

There is also rationale for testing dual-purpose interventions, in which IBS treatments may simultaneously reduce dementia risk. Precedent exists in other chronic conditions, including heart failure and diabetes, where disease-directed interventions have demonstrated concurrent reductions in dementia risk [4, 70]. For IBS specifically, probiotics have been evaluated as symptom relievers [71], although their effect on dementia risk remains unclear [72]. Interventions targeting gut health through probiotics, dietary modifications, or anti-inflammatory therapies could provide novel strategies for dementia prevention.

5. Conclusion

This meta-analysis provides the first synthesized evidence on the association between IBS and dementia, drawing on data from over 750,000 participants. Across analyses, IBS was consistently associated with statistically significantly higher odds of all-cause dementia and AD. These findings position IBS within a growing body of literature linking gastrointestinal disorders to neurodegenerative outcomes through shared gut-brain axis pathways.

However, the evidence remains observational, and the contributions of various types of bias cannot be excluded. IBS should not be regarded as an established causal risk factor for dementia based on current evidence, and no changes to dementia-specific clinical management are justified at this stage.

Well-designed prospective studies with IBS-specific exposure definitions, standardized dementia ascertainment, and rigorous control for confounders are needed to confirm and extend these findings. Mechanistic investigation of the gut-brain axis in IBS populations and subtype-specific dementia analyses represent equally important priorities. Finally, if the association is substantiated by future research, therapeutics targeting IBS pathophysiology may warrant evaluation as dual-purpose interventions, simultaneously managing IBS symptoms and potentially attenuating dementia risk.

Conflicts of Interest

They declare that they have no conflicts of interest to disclose.

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Large Language Model

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Authors Contribution

MMH contributed to conceptualization, software, writing review and editing, visualization, writing the original draft, data curation, formal analysis, resources, and methodology. NA and JL contributed to conceptualization, software, validation, supervision, writing review and editing, visualization, writing the original draft, data curation, formal analysis, resources, project administration, and methodology. SJ and LW contributed to software, writing review and editing, visualization, writing the original draft, data curation, resources, and methodology. All authors read and approved the final manuscript.

Data Availability

The data generated in this study are available upon request from the corresponding author.

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